

第6章 數學附錄

一 隨機變數線性函數的平均數與變異數。

設 X 為一隨機變數，其平均數為 μ_X ，變異數 σ_X^2 ，令 $Y = a \pm bX$ ，則 $E(Y) = \mu_Y = a \pm b\mu_X$ ， $V(Y) = \sigma_Y^2 = b^2\sigma_X^2$ ， $\sigma^2 = |b|\sigma_X$ 。

證明 Y 的期望值為：

$$E(Y) = E(a \pm bX) = E(a) \pm E(bX) = a \pm bE(X) = a \pm b\mu_X$$

變異數為：

$$\begin{aligned} V(Y) &= E[(a \pm bX) - E(a \pm bX)]^2 = E[a \pm bX - (a \pm bE(X))]^2 \\ &= E[bX - b\mu_X]^2 = b^2 E(X - \mu_X)^2 = b^2 \sigma_X^2 \end{aligned}$$

標準差為： $\sigma_Y = \sqrt{V(Y)} = |b|\sigma_X$

二 隨機變數 X 的變異數 $E(X - \mu)^2 = E(X^2) - [E(X)]^2$

證明

$$\begin{aligned} E(X - \mu)^2 &= \sum_{i=1}^n (x_i - \mu)^2 \cdot f(x_i) = \sum_{i=1}^n (x_i^2 - 2x_i\mu + \mu^2) \cdot f(x_i) \\ &= \sum_{i=1}^n [x_i^2 f(x_i) - 2\mu x_i f(x_i) + \mu^2 f(x_i)] \\ &= \sum_{i=1}^n x_i^2 f(x_i) - 2\mu^2 + \mu^2 = E(X^2) - \mu^2 = E(X^2) - [E(X)]^2 \end{aligned}$$

三 隨機變數 X 的偏度與峰度

證明

$$\alpha_3 = \frac{E(X - \mu)^3}{(\sqrt{E(X - \mu)^2})^3}, \quad \alpha_4 = \frac{E(X - \mu)^4}{[E(X - \mu)^2]^2}$$

式中： $E(X - \mu)^2 = E(X^2) - [E(X)]^2$ ， $E(X - \mu)^3 = E(X^3) - 3E(X)^2 E(X) + 2[E(X)]^3$

$E(X - \mu)^4 = E(X^4) - 4E(X^3)E(X) + 6E(X)^2[E(X)]^2 - 3[E(X)]^4 + 2[E(X)]^3 E(X^r)$ 稱為隨機變數 r 級原動差， $E(X - \mu)^r$ 稱為 r 級平均數動差。

四 二項機率分配之期望值與變異數

設 X 為一二項隨機變數，其機率函數為：

$$f(x) = C_x^n p^x q^{n-x} \quad x = 0, 1, 2, \dots, n$$

則其期望值為：

$$E(X) = np$$

變異數為：

$$V(X) = npq$$

證明 茲證明二項分配的期望值與變異數如下：

(1)期望值

$$\begin{aligned}
 E(X) &= \sum_{x=0}^n xf(x) = \sum_{x=0}^n x C_x^n p^x q^{n-x} = \sum_{x=0}^n x \frac{n!}{(n-x)!x!} p^x q^{n-x} \\
 &= \sum_{x=1}^n \frac{n(n-1)!}{(n-x)!(x-1)!} p^x q^{n-x} = np \sum_{y=0}^{n-1} \frac{(n-1)!}{(n-1-y)!y!} p^y q^{n-1-y} \\
 &\quad \text{「式中 } (y=x-1) \text{ 」} \\
 &= np(p+q)^{n-1} \\
 &\quad \text{(利用 } (n-1) \text{ 次方的二項展開式)} \\
 &= np
 \end{aligned}$$

(2)變異數

$$\begin{aligned}
 V(X) &= \sum_{x=0}^n (x - \mu)^2 f(x) = \sum_{x=0}^n x^2 f(x) - \mu^2 \\
 &= \sum_{x=0}^n x(x-1)f(x) + \sum_{x=0}^n xf(x) - \mu^2
 \end{aligned} \tag{①}$$

其中

$$\begin{aligned}
 \sum_{x=0}^n x(x-1)f(x) &= \sum_{x=0}^n x(x-1) C_x^n p^x q^{n-x} = \sum_{x=0}^n x(x-1) \frac{n!}{(n-x)!x!} p^x q^{n-x} \\
 &= n(n-1)p^2 \sum_{x=2}^{n-2} \frac{(n-2)!}{(n-2-x)!(x-2)!} p^{x-2} q^{n-x} = n(n-1)p^2(p+q)^{n-2}
 \end{aligned}$$

(利用 $(n-2)$ 次方的二次展開式)

$$= n(n-1)p^2 \tag{②}$$

將②代入①可得：

$$V(X) = n(n-1)p^2 + np - (np)^2 = -np^2 + np = np(1-p) = npq$$

五 超幾何機率分配之期望值與變異數

設 X 為一超幾何分配，則其期望值與變異數分別為：

$$E(X) = n \cdot \frac{K}{N}, \quad V(X) = n \cdot \frac{K}{N} \cdot \frac{N-K}{N} \cdot \frac{N-n}{N-1}$$

證明 分別證明如下：

(1)期望值

$$\begin{aligned}
 E(X) &= \sum_{x=0}^n x \frac{C_x^K C_{n-x}^{N-K}}{C_n^N} = \sum_{x=1}^n x \left(\frac{N!}{n!(N-n)!} \right)^{-1} \frac{K!}{x!(K-x)!} C_{n-x}^{N-K} \\
 &= \sum_{x=1}^n \left(\frac{N(N-1)!}{n(n-1)!(N-n)!} \right)^{-1} \cdot \frac{K(K-1)!}{(x-1)!(K-x)!} \cdot C_{n-x}^{N-K} \\
 &= n \frac{K}{N} \sum_{x=1}^n \frac{\frac{(x-1)!(K-x)!}{(N-1)!}}{\frac{(n-1)!(N-n)!}{(n-1)!(N-n)!}}
 \end{aligned}$$

$$\begin{aligned}
&= n \frac{K}{N} \sum_{x=1}^n \frac{C_x^{K-1} C_{n-x}^{N-K}}{C_{n-1}^{N-1}} = n \frac{K}{N} \sum_{y=0}^{n-1} \frac{C_y^{K-1} C_{n-1-y}^{N-K}}{C_{n-1}^{N-1}} \\
&\quad (x-1=y) \\
&= n \frac{K}{N} \frac{C_{n-1}^{N-1}}{C_{n-1}^{N-1}} \quad (\text{因為 } \sum_{i=0}^n \binom{a}{i} \binom{b}{n-i} = \binom{a+b}{n}) \\
&= n \frac{K}{N}
\end{aligned}$$

(2) 變異數

$$V(X) \text{ 可表為 : } V(X) = E[X(X-1)] + E(X) - [E(X)]^2 \quad \textcircled{1}$$

其中

$$\begin{aligned}
E[X(X-1)] &= \sum_{x=0}^n x(x-1) \frac{C_x^K C_{n-x}^{N-K}}{C_n^N} \\
&= n(n-1) \frac{K(K-1)}{N(N-1)} \sum_{x=2}^n \frac{C_{x-2}^{K-2} C_{n-x}^{N-K}}{C_{n-2}^{N-2}} \\
&= n(n-1) \frac{K(K-1)}{N(N-1)} \sum_{y=0}^{n-2} \frac{C_y^{K-2} C_{n-2-y}^{N-K}}{C_{n-2}^{N-2}} \\
&\quad (\text{註 } x-2=y) \\
&= n(n-1) \frac{K(K-1)}{N(N-1)} \quad \textcircled{2}
\end{aligned}$$

將 \textcircled{2} 代入 \textcircled{1} 可得

$$\begin{aligned}
V(X) &= E[X(X-1)] + E(X) - [E(X)]^2 \\
&= n(n-1) \frac{K(K-1)}{N(N-1)} + n \frac{K}{N} - [n \frac{K}{N}]^2 = n \frac{K}{N} \frac{N-K}{N} \frac{N-n}{N-1}
\end{aligned}$$

六 證明當 $N \rightarrow \infty$, 超幾何分配趨近於二項分配

$$\lim_{N \rightarrow \infty} \frac{C_x^K C_{n-x}^{N-K}}{C_n^N} = C_x^n p^x q^{n-x}$$

式中 : $\begin{cases} \lim_{N \rightarrow \infty} \frac{K}{N} = p \\ \lim_{N \rightarrow \infty} \frac{N-K}{N} = 1-p = q \end{cases}$

證明

$$\lim_{N \rightarrow \infty} \frac{C_x^K C_{n-x}^{N-K}}{C_n^N} = \lim_{N \rightarrow \infty} \frac{1}{N(N-1)\dots(N-n+1)} \cdot \frac{K(K-1)\dots(K-x+1)}{x!} \cdot \frac{n!}{n!}$$

$$\begin{aligned}
& \cdot \left(\frac{1}{(n-x)!} (N-K)(N-K-1)\dots(N-K-n+x+1) \right) \\
& = \lim_{N \rightarrow \infty} \frac{n!}{x!(n-x)!} \cdot \frac{1}{N(N-1)\dots(N-n+1)} \\
& \quad \cdot [K\dots(K-x+1)(N-K)\dots(N-K-n+x+1)]
\end{aligned}$$

(分子分母同除以 N^n)

$$\begin{aligned}
& = \lim_{N \rightarrow \infty} \frac{n!}{x!(n-x)!} \cdot \frac{1}{1 \cdot \frac{(N-1)}{N} \dots \frac{N-n+1}{N}} \\
& \quad \cdot \left[\frac{K}{N} \left(\frac{K-1}{N} \right) \dots \left(\frac{K-x+1}{N} \right) \left(\frac{N-K}{N} \right) \dots \left(\frac{N-K-n+x+1}{N} \right) \right] \\
& = \frac{n!}{x!(n-x)!} \cdot \left(\frac{K}{N} \right)^x \left(\frac{N-K}{N} \right)^{n-x} = C_x^n p^x q^{n-x}
\end{aligned}$$

因為

$$\begin{aligned}
& = \lim_{N \rightarrow \infty} \frac{1}{1 \cdot \frac{(N-1)}{N} \Lambda \frac{N-n+1}{N}} \cdot \frac{K}{N} \left(\frac{K-1}{N} \right) \Lambda \left(\frac{K-x+1}{N} \right) \left(\frac{N-K}{N} \right) \\
& \quad \Lambda \left(\frac{N-K-n+x+1}{N} \right) \\
& = \frac{1}{1 \cdot \left(1 - \frac{1}{N} \right) \dots \left(1 - \frac{n-1}{N} \right)} \cdot \frac{K}{N} \left(\frac{K}{N} - \frac{1}{N} \right) \Lambda \left(\frac{K}{N} - \frac{x-1}{N} \right) \left(\frac{N}{N} - \frac{K}{N} \right) \\
& \quad \Lambda \left(\frac{N-K}{N} - \frac{n-x-1}{N} \right)
\end{aligned}$$

但 $\lim_{N \rightarrow \infty} \frac{P}{N} = 0 \quad \forall P = 1, 2, \dots, (n-1)$

$$= \frac{\frac{K}{N} \Lambda \frac{K}{N} \cdot \frac{N-K}{N} \Lambda \frac{N-K}{N}}{1 \cdot 1 \Lambda 1}$$

(因 $\frac{K}{N}$ 有 x 個, $\frac{N-K}{N}$ 有 $n-x$ 個)

$$= \left(\frac{K}{N} \right)^x \left(\frac{N-K}{N} \right)^{n-x}$$

七 證明泊松機率分配之期望值與變異數。

設 X 為一泊松分配，則其期望值與變異數為：

$$E(X) = \lambda \quad V(X) = \lambda$$

證明 證明如下：

(1) 期望值

$$E(X) = \sum_{x=0}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x!} = \lambda \sum_{x=1}^{\infty} \frac{e^{-\lambda} \lambda^{x-1}}{(x-1)!} = \lambda e^{-\lambda} e^{\lambda} = \lambda$$

(因 $\sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!} = e^{\lambda}$)

(2) 變異數

$$\begin{aligned} V(X) &= E[X(X-1)] + E(X) - [E(X)]^2 \\ &= \sum_{x=0}^{\infty} x(x-1) \frac{e^{-\lambda} \lambda^x}{x!} + \lambda - \lambda^2 = \lambda^2 \sum_{x=2}^{\infty} \frac{e^{-\lambda} \lambda^{x-2}}{(x-2)!} + \lambda - \lambda^2 \\ &\quad (\text{因 } \sum_{x=2}^{\infty} \frac{\lambda^{x-2}}{(x-2)!} = e^{\lambda}) \\ &= \lambda^2 + \lambda - \lambda^2 = \lambda \end{aligned}$$

八 證明當 $n \rightarrow \infty$, np 固定, 二項分配趨近於泊松分配

$$\lim_{n \rightarrow \infty} C_x^n p^x q^{n-x} = \frac{\lambda^x e^{-\lambda}}{x!}$$

式中: $\lambda = np$

證明 證明如下:

$$\begin{aligned} \lim_{n \rightarrow \infty} C_x^n p^x q^{n-x} &= \lim_{n \rightarrow \infty} \frac{n!}{x!(n-x)!} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x} \\ &= \lim_{n \rightarrow \infty} \frac{\lambda^x}{x!} \cdot \frac{n(n-1)...(n-x+1)}{n^x} \cdot \left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-x} = \frac{\lambda^x e^{-\lambda}}{x!} \dots (\lambda = np) \end{aligned}$$

因為

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{n(n-1)...(n-x+1)}{n^x} &= \lim_{n \rightarrow \infty} \left[\left(1 - \frac{1}{n}\right) \left(1 - \frac{2}{n}\right) ... \left(1 - \frac{x-1}{n}\right) \right] = 1 \\ \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n &= e^{-\lambda} \quad \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{-x} = 1 \end{aligned}$$

九 證明幾何分配的期望值與變異數

$$E(X) = \frac{1}{p}, \quad V(X) = \frac{1-p}{p^2}$$

(1) 期望值

$$E(X) = \sum_{x=1}^{\infty} x q^{x-1} p = p \sum_{x=1}^{\infty} x \cdot q^{x-1} \quad \textcircled{1}$$

令

$$T = \sum_{x=1}^{\infty} x \cdot q^{x-1} = 1 + 2q + 3q^2 + \Lambda \quad \textcircled{2}$$

②乘上 q 可得:

$$qT = \sum_{x=1}^{\infty} x \cdot q^x = q + 2q^2 + 3q^3 + \Lambda \quad \textcircled{3}$$

② - ③式可得:

$$T - qT = T(1 - q) \\ = (1 + q + q^2 + \dots) = \frac{1}{1 - q} \quad (\text{無窮項等比級數公式}) \quad ④$$

由④可解得：

$$T = \frac{1}{(1 - q)^2} \quad ⑤$$

⑤代入①可得

$$E(X) = p \cdot \frac{1}{(1 - q)^2} = p \cdot \frac{1}{p^2} = \frac{1}{p}$$

(2) 變異數

$$V(X) = E(X^2) - [E(X)]^2 = E[X(X - 1)] + E(X) - [E(X)]^2$$

$$= \frac{2q}{p^2} + \frac{1}{p} - \frac{1}{p^2} = \frac{q}{p^2} = \frac{1-p}{p^2}$$

$$\text{因為 } \frac{dq^x}{dq^2} = x(x-1)q^{x-2}$$

$$\text{而 } E[X(X - 1)] = \sum_{x=1}^{\infty} x(x-1)q^{x-1} p$$

$$= pq \sum_{x=1}^{\infty} x(x-1)q^{x-2} = pq \frac{d^2}{dq^2} \left(\sum_{x=0}^{\infty} q^x \right) = pq \frac{2}{(1-q)^2} = \frac{2}{p^2}$$

其中

$$\sum_{x=0}^{\infty} q^x = \frac{1}{1-q} \quad \frac{d^2}{dq^2} \left(\frac{1}{1-q} \right) = \frac{2}{(1-q)^3}$$